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# Discrete-Time Analysis: Deriving the Distribution of the Number of Events in an Arbitrarily Distributed Interval 

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#### Abstract

For the context of discrete-time analysis, this document describes how to derive the distribution of the number of events in an arbitrarily distributed interval.


Keywords: Discrete-Time Analysis.

## 1 Introduction

For the context of discrete-time analysis, this document describes how to derive the distribution of the number of events in an arbitrarily distributed interval. All material is based on [1].

In this work, we use the following notation to distinguish between random variables (RVs), their distributions, and their distribution functions. An RV is represented by an uppercase letter, e.g., $X$. The distribution of $X$ is denoted by $x(k)$ and is defined as

$$
x(k)=P(X=k),-\infty<k<\infty .
$$

Furthermore, the distribution function of $X$ is written as $X(k)$ and is defined as

$$
X(k)=\sum_{i=0}^{k} x(i),-\infty<k<\infty
$$

Finally, $\mathrm{E}[X]$ denotes the mean of $X$ and $*$ refers to the discrete convolution operation, i.e.,

$$
a_{3}(k)=a_{1}(k) * a_{2}(k)=\sum_{j=-\infty}^{\infty} a_{1}(k-j) \cdot a_{2}(j)
$$

## 2 Process

We consider a time discrete renewal process whose interarrival times are defined by the RV $A$. The process is observed during an interval of length $T$ that follows the distribution $\tau(m), m=0,1, \ldots$. We assume the two endpoints of the interval to be located immediately before discrete points in time (cf. Figure 1). Furthermore, $A^{*}$ refers to the forward recurrence time of RV $A$, i.e., the time between any given time and the arrival of the next event. The distribution $x(j)$ of the number of events $X$ during a random observation interval is derived in this section.


Figure 1: Number of events in a random interval.

Using the law of total probability, the distribution $x(j)$ can be expressed as follows.

$$
\begin{equation*}
x(j)=\sum_{m=0}^{\infty} \mathrm{P}(X=j \mid T=m) \cdot \mathrm{P}(T=m)=\sum_{m=0}^{\infty} x(j \mid m) \tau(m) . \tag{1}
\end{equation*}
$$

Let $F^{(j)}$ denote the RV for the time between the start of the observation interval and the $j$-th event (cf. Figure 1). The following two equations describe this RV and its distribution formally.

$$
\begin{align*}
& F^{(j)}=A^{*}+\underbrace{A+\cdots+A}_{(j-1) \text { times }} .  \tag{2}\\
& f^{(j)}(k)=a^{*}(k) * \underbrace{a(k) * \cdots * a(k)}_{(j-1) \text { times }} . \tag{3}
\end{align*}
$$

Hence, the conditional probability $x(j \mid m)$ on the right hand side of Equation 1 calculates as follows.

$$
\begin{equation*}
x(j \mid m)=\mathrm{P}\left(F^{(j)}<m \leq F^{(j+1)}\right)=\mathrm{P}\left(F^{(j)}<m\right)-\mathrm{P}\left(F^{(j+1)}<m\right) . \tag{4}
\end{equation*}
$$

Taking into account the special case that an interval $T$ with length $m=0$ does not contain any events, we get:

$$
\begin{align*}
& x(j \mid 0)=\delta(j)=\left\{\begin{array}{ll}
1 & j=0 \\
0 & \text { otherwise }
\end{array}, \quad m=0\right. \\
& x(j \mid m)=\sum_{i=0}^{m-1}\left(f^{(j)}(i)-f^{(j+1)}(i)\right), \quad m=1,2, \ldots \tag{5}
\end{align*}
$$

Finally, Equations 1 and 5 are used to determine the distribution of the number of events during the observation interval $T$ :

$$
\begin{equation*}
x(j)=\tau(0) \delta(j)+\sum_{m=1}^{\infty} \tau(m) \sum_{i=0}^{m-1}\left(f^{(j)}(i)-f^{(j+1)}(i)\right), \quad j=0,1, \ldots \tag{6}
\end{equation*}
$$

## References

[1] P. Tran-Gia, "Zeitdiskrete Analyse verkehrstheoretischer Modelle in Rechner- und Kommunikationssystemen - 46. Bericht über verkehrstheoretische Arbeiten," 1988.

